# Nested sampling for small p-values

Andrew Fowlie, Sebastian Hoof, and Will Handley (May 2021). In: arXiv: 2105.13923 [physics.data-an] **Q** github.com/andrewfowlie/ns\_for\_p\_values

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- 1. Why small *p*-values?
- 2. Why is it difficult to compute small *p*-values?
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# Why small *p*-values?

#### **P**-value

The *p*-value, *p*, is the probability of observing data as or more extreme than that observed, given the null hypothesis,  $H_0$ , i.e.,

$$p = P(\lambda \ge \lambda_{\text{Observed}} \mid H_0)$$

where  $\lambda$  is a test-statistic that summarises the data and defines extremeness.

## **P**-values in high-energy physics



In high-energy physics, we want to discover new phenomena and new particles. Perform null hypothesis test:

- $H_0$  Standard Model (SM) backgrounds only
- $H_1 SM$  + new physics, e.g. Higgs boson or supersymmetric particles

We conventionally require a tiny global *p*-value less than about  $10^{-7}$ , corresponding to  $5\sigma$ 

#### Classic example. Higgs discovery in 2012 (Aad et al. 2012).



Wait until reach  $5\sigma$  global. We need to compute tiny *p*-values.

## Why is it difficult to compute small *p*-values?

#### Illustrate problem with two-dimensional data, x.



In reality, red region exponentially tiny. Illustrate with squares but assume nothing about geometry/topology in problem or solution.

#### Sample from whole sample space – Monte Carlo



Draw *n* samples from whole sampling distribution. We really need at least one sample to fall in red region.

Estimate *p* by fraction of them that fall in red region

$$\hat{p} = \frac{m}{n}$$

Error of order Wald (Brown, Cai, and DasGupta 2001)

$$\frac{\Delta p}{p} = \sqrt{\frac{1/p}{n}}$$

Usual  $1/\sqrt{n}$  statistical error scaling. For fixed fractional uncertainty, number of samples scales as 1/p

*Need*  $n \gtrsim 1/p$  *at very least for reasonable estimate.* 

If we have some more information, we can compute *p* analytically.

- E.g., I know the red area is a box of side 0.05.  $p = 0.05^2$ .
- Sometimes, we know (or hope!) our problem satisfies certain regularity conditions and large sample limit. We can apply asymptotics (Cowan et al. 2011).

#### But the conditions aren't always satisfied.

We want generality. Generality is power to tackle any problem we want.

# **Nested sampling**

Nested sampling (Skilling 2004; Skilling 2006) originally algorithm for Bayesian computation. We reinterpreted it in the context of *p*-values

- Evolves collection of *n*live live points to greater and greater test-statistics
- Evolution controlled by single user parameter  $n_{live}$
- Replaces one point at a time
- Meta-algorithm different solutions to finding replacement points
- Many existing public implementations

- 1. Sample *n* points from the sampling distribution
- 2. Rank them by test-statistic
- 3. Delete the least extreme half

You just compressed by factor 1/2! Repeat it i times and you'll achieve exponential compression  $1/2^i$  at a constant rate!

If it took  $n_{\text{iter}}$  iterations to reach the area corresponding to p-value, then

$$p=\frac{1}{2^{n_{\text{iter}}}}$$

## **Complete algorithm**

- Draw n<sub>live</sub> sets of pseudo-data from the sampling distribution — the live points
- 2. Initialize  $n_{\text{iter}} = 0$
- 3. repeat

4. 
$$n_{\text{iter}} = n_{\text{iter}} + 1$$

- 5. Find the minimum TS  $\lambda^*$  amongst the live points
- 6. Replace live point corresponding to  $\lambda^*$  by one drawn from the sampling distribution subject to  $\lambda > \lambda^*$

7. until 
$$\lambda^{\star} \geq \lambda_{Observed}$$

8. return *Estimate of*  $p = e^{-n_{iter}/n_{live}}$ 

- By replacing half the points, we compress by 1/2 per iteration
- That isn't optimal better to replace one point at a time
- Compress by about  $e^{-1/n_{\text{live}}} \simeq 1 1/n_{\text{live}}$  each iteration

This breaks the computation of a tiny p-value into the product of  $n_{iter}$  moderate factors

$$p = \prod_{i=1}^{n_{iter}} e^{-1/n_{live}} = e^{-n_{iter}/n_{live}}$$



Compress into tail of TS distribution. Stop once we get to *p*-value.

For fixed fractional uncertainty on *p*, we expect to obtain a speed-up versus MC

$$\frac{\text{Evaluations for NS}}{\text{Evaluations for MC}} = \frac{(\log^2 1/p)/\epsilon}{1/p}$$

Massive gains for small p! Provided that the efficiency factor  $\epsilon$  doesn't spoil things.

# **Exploration**

So far nothing depended on the problem at hand or the dimension of the sampling space. I glossed over a detail though

- How to draw a replacement live point from the constrained sampling distribution, λ > λ\*?
- Involves trial and error, and some inefficiency (the factor  $\epsilon$ )
- Re-introduces dependence on dimensionality of the sampling space (though it needn't be exponential)

### **Strategies**

MultiNest (Feroz and Hobson 2008; Feroz, Hobson, and Bridges 2009; Feroz et al. 2013) — bound live points by ellipsoids that approximate the  $\lambda^*$  contour and sample from them

**PolyChord** (Handley, Hobson, and Lasenby 2015a; Handley, Hobson, and Lasenby 2015b). Take walk starting from a randomly chosen existing live point — in **PolyChord** a slice sampling walk.



## **Example & Tools**

The *p*-value associated with *d* independent Gaussian measurements

- *d* dimensional sampling space,  $\mathbf{x}_i \sim \mathcal{N}(\mu, \sigma^2)$ .
- Test-statistic

$$\lambda = \sum_{i=1}^{d} \left( \frac{\mathbf{x}_i - \mu}{\sigma} \right)^2$$

• We know analytically

$$\lambda \sim \chi_d^2$$
 such that  $p = 1 - F_{\chi_d^2}(\lambda_{\text{Observed}})$ 

• Toy example that allows us to easily control dimension, size of *p* and check correctness

#### Number of evaluations for fixed fractional uncertainty



Perfect NS means if 100% efficiency,  $\epsilon = 1$ , was possible

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- For large *p*, nested sampling could only narrowly beat MC in the best-case scenario, and typically worse. MC just fine when *p* is moderate
- For small  $p \lesssim 4\sigma$ , the scaling kicks in. Nested sampling wins by orders of magnitude
- Nested sampling performance depends on dimensionality but even for 30*d* sampling space, winning by  $10^6$  at  $7\sigma$

Import our PolyChord wrapper from our code
O github.com/andrewfowlie/ns\_for\_p\_values

>>> from p\_value import pc

This wraps PolyChord to enable the required stopping conditions and return the *p*-value from NS.

We also wrapped dynesty and MultiNest. Healthy ecosystem of publicly available NS implementations and NS analysis software.

Make a function that transforms U(0, 1) draws to chi-squared draws. These chi-squared draws are the pseudo-data

>>> from scipy.stats import chi2

>>> def transform(unit\_hyper\_cube): ... return chi2.ppf(unit\_hyper\_cube, df=1)

This is the inverse transform method using the chi-squared distribution's percent point function — chi2.ppf

# Define the test-statistic — here it's just the sum of the chi-squared draws

>>> def test\_statistic(data):
... return data.sum()

Run PolyChord on this problem for 5 draws and an observed  $\lambda = 50$ , and using 100 live points.



Same signature for our MultiNest and dynesty wrappers, just replace pc with mn or dynesty.

Simple easy-to-use signature. Wraps established, fast libraries for NS.

The results are displayed as

```
p-value = (1.5910 +/- 0.7161)e-09.
log10(p-value) = -8.7983 +/- 0.19548.
Signifiance = 5.922 sigma.
Function calls = 309765
```

We computed a *p*-value of 1 in a billion in only 300,000 calls! Agrees with analytic result within uncertainty:

>>> chi2.sf(50, df=5)
1.3857973367009573e-09

- Nested sampling particularly suitable for *p*-value computation, as it naturally builds path to the *p*-value
- Orders of magnitude faster than Monte Carlo for small p, as scaling log<sup>2</sup> 1/p rather than 1/p for fixed relative error
- Performance understood theoretically and demonstrated numerically
- With small modifications, standard nested sampling software can be used to compute *p*-values
- **O** github.com/andrewfowlie/ns\_for\_p\_values

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